

1-dimensional Gaussian distribution $N(\mu, \sigma)$

$$\frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2}$$

2-D Gaussian p. 191

$$f_{X,Y}^{\sim}(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} e$$

$$-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_1}{\sigma_1}\right)^2 + \left(\frac{y-\mu_2}{\sigma_2}\right)^2 - \dots \right]$$

$$\dots - \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} \Bigg]$$

$$= (\text{const}) e^{x^2, y^2, xy, x, y, \text{const}}$$

$$\underline{1-d)} \quad \text{Prob}(a < x < b) = \int_a^b f_x(x) dx$$

2-d)

$$\text{Prob}(a < x < b \text{ and } c < y < d) = \int_a^b \int_c^d f_{X,Y}(x,y) dy dx$$

Joint distribution function

$$\mathcal{E}(X) = \mu_1 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f_{X,Y}(x,y) dy dx$$

$$\mathcal{E}(Y) = \mu_2 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f_{X,Y}(x,y) dx dy$$

$$\mathcal{E}(X^2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x^2 f_{X,Y}(x,y) dx dy \dots$$

$$\sigma_1^2 = \mathcal{E}[(X - \mu_1)^2] = \mathcal{E}[X^2] - \mu_1^2$$

Marginal Distributions

Prob ($a < x < b$ and y is anything)

$$= \int_{x=a}^b \int_{y=-\infty}^{\infty} f_{X,Y}(x,y) dy dx$$

Marginal on X : $f_X(x) = \int_{y=-\infty}^{\infty} f_{X,Y}(x,y) dy$

Fact The marginal density function for
the 2-D Gaussian
is the 1-D Gaussian

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi} \sigma_1 \sigma_2} e^{-\frac{1}{2\pi \sigma_1^2 \sigma_2^2} \left[\left(\frac{x-\mu_1}{\sigma_1} \right)^2 + \left(\frac{y-\mu_2}{\sigma_2} \right)^2 - 2\rho \frac{(x-\mu_1)(y-\mu_2)}{\sigma_1 \sigma_2} \right]} dy$$
$$= \frac{1}{\sqrt{\pi} \sigma_1} e^{-\frac{1}{2} \left(\frac{x-\mu_1}{\sigma_1} \right)^2}$$

~~Conditional Prob~~

Incidental Remark :

You can write the bivariate form, in the 2-D exponential, in a matrix form :

$$- [x-\mu_1, y-\mu_2] \begin{bmatrix} X & X \\ X & X \end{bmatrix} \begin{bmatrix} x-\mu_1 \\ y-\mu_2 \end{bmatrix}$$

Conditional probability.

$$P(A, \text{ given } B)$$

If A & B are independent,
This is the same as P(A) given

Basic rule:

$$P(B)P(A|B) = P(A+B) = P(A)P(B|A)$$

$$P(B|A) = \frac{P(A+B)}{P(A)}$$

$$\text{conditional} = \frac{\text{joint}}{\text{marginal}}$$

$$P(B|A) = P(A|B) P(B) / P(A)$$

Bayes' Thm.

IF A & B are independent,

$$P(A+B) = P(A)P(B)$$

Fact: for density functions,
 conditional = $\frac{\text{joint}}{\text{marginal}}$

\mathbb{P}

$$\text{Prob}(x \leq X \leq x+dx, \text{ given that } y \leq Y \leq y+dy) = \frac{\text{Prob}(x \leq X \leq x+dx \text{ and } y \leq Y \leq y+dy)}{\text{Prob}(y \leq Y \leq y+dy)}$$

"Y=y"

conditional

$$f_{X|Y}(x; y) dx = \frac{f_{XY}(x, y) dx dy}{f_Y(y) dy}$$

corrected at end of lecture

$$f_{X|Y}(x; Y=y) dx$$

(Trust me)

$$f_{X|Y}(x,y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

$$f_{X|Y}(x; y) = \frac{f_{XY}(x, y)}{f_Y(y)}$$

For the Gaussian

LHS to be determined } will turn out to be Gaussian in x

p. 193, 4

$$N\left(\mu_1 + \rho \frac{\sigma_1}{\sigma_2} (y - \mu_2), \sigma_1 \sqrt{1 - \rho^2}\right)$$

mean

$$f_{X|Y}(x; y) = \frac{f_{XY}(x, y)}{f_Y(y)}$$

$$\frac{\frac{1}{\sqrt{2\pi}\sigma_1\sqrt{1-\rho^2}} e^{-\frac{1}{2} \frac{\left[\frac{x-\mu_1 - \rho \frac{\sigma_1}{\sigma_2} (y-\mu_2) \right]^2}{\sigma_1^2 (1-\rho^2)}}}{\int_{-\infty}^{\infty} \left(\text{above} \right) dx} = \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{1}{2} \left(\frac{y-\mu_2}{\sigma_2} \right)^2}$$

NOTE: If $\rho=0$ X & Y are independent

because

$$f_{XY} = \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{1}{2} \left(\frac{x-\mu_1}{\sigma_1} \right)^2} \cdot \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{1}{2} \left(\frac{y-\mu_2}{\sigma_2} \right)^2}$$